

# Image Denoising Using Standard BP Algorithm

T. Uma Mageswari

Assistant Professor (IT), Adhiparasakthi Engineering College, Melmaruvathur, India

**Abstract:** *From the perspective of the Bayesian approach, the denoising problem is essentially a prior probability modeling and estimation task. In this paper, we propose an approach that exploits a hidden Bayesian network, constructed from wavelet coefficients, to model the prior probability of the original image. Then, we use the belief propagation (BP) algorithm, which estimates a coefficient based on all the coefficients of an image, as the maximum-a-posterior (MAP) estimator to derive the denoised wavelet coefficients. We show that if the network is a spanning tree, the standard BP algorithm can perform MAP estimation efficiently. Our experiment results demonstrate that, in terms of the peak-signal-to-noise-ratio and perceptual quality, the proposed approach outperforms state-of-the-art algorithms on several images, particularly in the textured regions, with various amounts of white Gaussian noise.*

**Keywords:** Bayesian network, image denoising, wavelet transform

## 1. Introduction

COMPLEX phenomena usually involve a large number of hidden variables and data sources. Graphical models provide a unifying framework for modeling the probability distributions of such phenomena by decomposing joint probability distributions into a set of local constraints and dependencies [1]. After formulating a problem as a graphical model, a wide range of statistical learning and inference algorithms can be applied directly to derive a solution. Bayesian networks are probably the most popular type of (directed) graphical model. In this paper, our objective is to construct a Bayesian network from a single image for denoising purposes. To do this, we need to overcome two difficulties: 1) constructing a Bayesian network is computationally inefficient and 2) the data over-fitting problem, which exaggerates minor fluctuations in the input data.

The construction of a Bayesian network involves prior knowledge of the probability relationships between the variables of interest. Learning approaches are widely used to construct Bayesian networks that best represent the joint probabilities of training data [2] - [5]. In practice, an optimization process based on a heuristic search technique is used to find the best structure over the space of all possible networks.

However, the approach is computationally intractable because it must explore several combinations of dependent variables to derive an optimal Bayesian network. The difficulty is resolved in this paper by representing the data in wavelet domains and restricting the space of possible networks by using certain techniques, such as the "maximal weighted spanning tree" (MWST). Three wavelet properties—sparsity, clustering, and persistence—can be exploited to reduce the computational complexity of learning Bayesian network. First, the wavelet transform of a natural image tends to be sparse with large coefficients at the edges. The sparsity reduces the number of variables required to construct a graph. Second, the adjacent wavelet coefficients tend to have similar values as a cluster. Third, wavelet coefficients at the same location and orientation tend to be positively correlated in adjacent scales.

The over-fitting problem occurs because the underlying network is too complex; for example, there may be too many parameters for the number of observations [6]. In the effort of using hidden Markov tree (HMT) model [7]–[9] to capture the joint statistics of wavelet coefficients, the marginal probability of each wavelet coefficient is modeled as a mixed density function with a hidden state variable; for example, an  $M$ -state Gaussian mixture model for a wavelet coefficient consists of  $M$  states and  $M$  Gaussian conditional probability density functions (pdf), one for each state.

To generate wavelet coefficients, the HMT first draws a state values based on the pdf  $p(s)$ , and then draws an observation according to the conditional probability  $f_M^s$  of the state. The pdf of the wavelet coefficient is given by  $p(s)f_M^s(\cdot|s)$ , where the conditional pdf  $f_M^s(\cdot|s)$  is written as a parametric formula that depends on certain parameters. Because several parameters are used to estimate a wavelet coefficient, the HMT approach can only be used to model the marginal probability when the number of training images is large; then the underlying parameters can be estimated accurately. If there is only one image, the over-fitting problem under the HMT approach would be severe. Thus, instead of associating each wavelet coefficient with a random variable, we split all the wavelet coefficients into equal-size blocks and assume that the blocks are independently sampled from a matrix of random variables (called a wavelet patch). The approach allows us to estimate non-parametric statistics, which do not require a pdf assumption, from the samples in the wavelet patch.

The image denoising problem is particularly serious in modern image capturing devices because the increase in the sensor's density per unit area of a chip reduces the signal-to-noise (SNR) and increases the capturing device's sensitivity to noise [10]. The state-of-the-art denoising algorithms are based on the non-local means approach [11]–[13], which exploits the self-similarity and redundancy in an image. The most representative approach is the block matching and 3-D filtering (BM3D) algorithm [13]. It combines similar 2-D patches that can be overlapped to form a 3-D group, and then uses collaborative 3-D filtering techniques to perform non-local filtering. The filtered blocks are returned to their original positions, and the final estimate of a pixel is

computed as the weighted average of the estimates of the pixel in several different blocks. This simple approach is very efficient and it yields better results than regular non-local means approaches. In [14], a denoising Wiener filter, motivated by the statistical analysis of the performance bounds of patch-based methods, is proposed. The filter's parameters are estimated from geometrically as well as photo metrically similar patches. Recent developments in sparse representation have been used together with the non-local means approach for denoising purposes [15], [16]. The sparse model approach assumes that image patches can be represented sparsely by an over-complete redundant dictionary, which can be learned from a family of training data. In [17], Milanfar demonstrates that the non-local means approach as well as a recent multi-dimensional filtering approaches (e.g., bilateral filtering and kernel regression) are operations that are adaptive to the local structure of an image. He also presents a general framework for understanding the basic principles behind the approaches.

The Bayesian approach is also widely used to resolve the image denoising problem. The Bayesian formula indicates that the denoising problem is essentially a prior probability modeling and estimation task. If  $y = x + n$ , where  $n$  is white Gaussian noise with known variance, then the Bayesian formula is  $P(x|y) = P(y|x)P(x)$ , where  $P(y|x) = P_n(y-x)$  is the noise probability.

The maximum-a-posterior (MAP) solution of  $P(x|y)$  is determined by the prior probability  $P(x)$ . The structure of image prior is usually modeled by Markov Random Fields (MRFs), where the probability of a pixel depends solely on the joint probability of the pixels in its neighborhood. According to the Hammersley-Clifford theorem, the probability distribution of an MRF is the Gibbs distribution whose energy function is the sum of the potential functions defined on the cliques (i.e., maximal complete subgraphs) in image neighborhoods.

Many wavelet-based denoising algorithms integrate the wavelet properties in MRF structure in the structure of a denoised image. The main differences between the algorithms are the methods used for neighborhood selection, the modeling of the original image over the neighborhood, and the techniques employed to derive solutions. In [18], the estimated wavelet coefficient at index  $l$  (position, scale),  $w_l$ , is obtained by

$$\hat{w}_l = w_l P(x_l|M) \approx w_l P(M|x_l) P(x_l),$$

where  $w_l$  is the observed wavelet coefficient,  $x_l$  is the hidden label,  $P(x_l|M)$  is the probability of the label  $l$  as a node, and  $M$  is the measurement derived by inter-scale Lipschitz exponent estimation [19], [20].

The prior probability  $P(x_l)$  at the scale of  $l$  is modeled by a  $2 \times 2$  MRF, where the potential functions between two neighboring wavelet coefficients on hand-chosen cliques are defined as  $x_i x_j$ , with each variable taking the value 1 or -1. The approach is extended in [21] by incorporating robust

inter-scale estimators  $M$  and  $P(x_l|M)$  and generalized an isotropic MRF prior  $P(x_l)$  on each scale.

Simoncelli examined the empirical statistical properties of images in adjacent scales and presented an inter-scale probability model for the wavelet coefficients in two adjacent scales [22]. The joint statistical model assumes that the density of an estimated wavelet coefficient is conditionally Gaussian, where the variance is a linear combination of the squared coefficients in a local neighborhood.

The neighborhood consists of coefficients at other orientations and adjacent scales, as well as adjacent spatial locations. For example, Simoncelli's model uses a neighborhood comprised of the 12 nearest spatial neighbors in the same sub band, the 5 nearest coefficients in sub bands at other orientations on the same scale, the 9 nearest coefficients in the adjacent subband of the coarser scale, and some coefficients in other subbands.

The BLS-GSM algorithm [23] models the distribution of a vector of wavelet coefficients in a  $3 \times 3$  region, together with the coefficient at the center location and the same orientation at the next coarser scale, as a Gaussian Scale Mixture (GSM). Then, the Bayesian least square method is used to estimate the wavelet coefficient at the center of the neighborhood system. In [24], a mixture of Gaussian Scale Mixtures (MGSM) is proposed to make the GSM model adaptive to the image content. Its denoising performance is almost as good as that of the BM3D algorithm. In [25], a dimension reduction technique is applied to the MGSM to reduce the computational cost and avoid the curse of dimensionality problem in learning the high number of free parameters of the model.

The MRF is modeled on a non-directed graph. In this paper, we propose an approach that uses a hidden directed graph to model the prior probability of an image. Specifically, the graph is a Bayesian network with a multi-layer network structure constructed from the wavelet coefficients of an image. Our approach has two advantages over existing approaches.

First, the MAP solution can be derived by using the standard BP algorithm [26], [27]. BP inference passes messages forward from coefficients at coarser scales to finer scales and back-ward from finer scales to coarser scales. Forward message passing tends to smooth fluctuations in fine scales, while backward message passing tends to retain the fine details of an image.

Thus, our method is more capable of retaining the fine structure of an image than existing approaches. The second advantage is that the hidden structure is derived by a data-adaptive process.

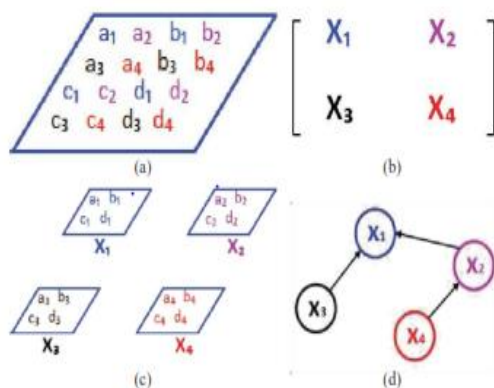
In addition, the prior probabilities over inter-scale edges and intra-scale edges are modeled in a similar way to those in [22] and [28] respectively.

We also analyze the complexity of estimating the solution of a Bayesian network and show that BP inference can derive

the MAP solution efficiently provided that the Bayesian network is a spanning tree.

To evaluate our approach, we compare its performance with that of other approaches, including BM 3D, and demonstrate that it yields a better peak-signal-to-noise ratio (PSNR) as well as better perceptual quality on the textured are as of an image.

The remainder of this paper is organized as follows. In Section II, we explain the rationale behind the proposed data-adaptive graph modeling of an image. We also analyze the computational cost of using BP inference to estimate the MAP solution for the proposed approach. In Section III, we present the method used to construct and model wavelet Bayesian networks; and in Section IV, we describe how the networks are used for denoising. In Section V, we discuss the proposed denoising algorithm and compare its performance with that of other approaches. Section VI contains some concluding remarks.



**Figure 1:** Constructing a hidden directed graph from a wavelet subband. (a) Subband of  $4 \times 4$  coefficients. (b) Wavelet patch is a matrix of  $2 \times 2$  random variables,  $X_1$ ,  $X_2$ ,  $X_3$ , and  $X_4$ . (c) Subband is divided into four rectangles, each containing  $2 \times 2$  coefficients. The coefficients at the same location in each rectangle are assigned to the node at the same location in the wavelet patch, and the coefficients of the same color are grouped as the observations of a node. (d) Directed edges are constructed to form a spanning tree.

## 2. Data-Adaptive Hidden Network Approach

In this section, we describe the proposed framework for the constructing a data-adaptive  $\frac{m^2-1}{m^2}$  graph structure and formulating the prior  $m$  and probability of the original image in the wavelet domain. We also explain how the graph structure is used with the BP algorithm to derive the MAP solution.

We assume that the wavelet coefficients of the original image represent a realization of a hidden graph. To construct the graph, we first create a matrix of random variables (a wavelet patch) for each subband (indexed by orientation and scale), and assign the subband coefficients as the observed data of the random variables. Then, a subgraph (network) is constructed from the wavelet patch as follows.

First, we associate each random variable with a node; hence, the coefficients assigned to the random variable can be regarded as node's observed data. Second, the arcs (directed edges) in the subgraph are derived according to a data dependence measurement between the observed data in any pair of nodes.

In this way, we can construct a one-layer network from the wavelet subband and associate the sub band's coefficients with the nodes in the network. The procedure is shown in Fig.1. Finally, the one-layer networks of adjacent subbands can be linked by inter-scale arcs, from coarser scales to finer scales, to form a multi-layer network structure, as shown in Fig. 2(a)–(c).

Two problems may arise during the above construction procedure: 1) the coefficient and wavelet patch association problem, which involves associating subband coefficients with a wavelet patch; and 2) the graph selection problem, i.e., determining the type of graph to construct. To solve the first problem, we propose the following heuristic procedure. Assume that the wavelet patch is a matrix of  $m \times m$  random variables. Let the size of a subband be  $N \times N$  and let  $m$  divide  $N$ . We partition the subband into  $(\frac{N}{m})^2$  rectangular blocks, each of which contains  $m \times m$  coefficients. Then, the coefficient at location  $(i, j)$  in each block is assigned as a realization of the random variable at location  $(i, j)$  in the wavelet patch. Thus, each random variable has  $\frac{N^2}{m^2}$  sampled observations.

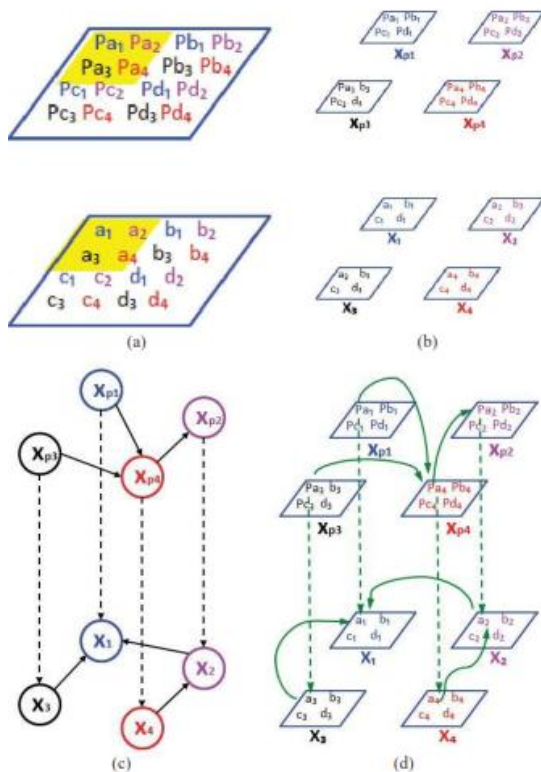
For the second problem, we analyze the computational cost of a graph structure for which the MAP solution can be derived efficiently by the standard BP algorithm. The standard implementation of the message passing algorithm in BP on  $m \times m$  cliques runs in  $O(N^2 k^{m \times m} T)$ , where  $N^2$  is the number of coefficients in a subband,  $k$  is the number of labels for each coefficient, and  $T$  is the number of iterations. Basically, computing each message takes  $O(k^{m \times m})$  time and there are  $O(N^2)$  messages per iteration. The computational cost of the conventional BP algorithm can be reduced if the algorithmic techniques in [29]–[31] are used for Bayesian inference.

If parallelism is also exploited as described in [32], BP inference can achieve a near linear parallel scaling performance. We use a graph  $G(V, E)$  to represent the (one-layer) subgraph structure, where  $V$  is the node set and  $E$  is the arc set. If  $G$  is a loopy Bayesian network, the BP algorithm sometimes yields surprisingly good approximate results; however, sometimes, it fails to produce any results, even after a large number of iterations. BP can derive an exact MAP solution in two iterations if  $G$  is a directed-acyclic-graph (DAG), but the inference cost depends on the structure of the DAG. A DAG that incurs a high inference cost can be constructed as follows: let the nodes in  $G$  be indexed from 1 to  $m^2$ . For each node  $j$ , passing at node  $j$  For typical values of  $k$  and  $m$ , which are 512 and 4 respectively, BP inference of the DAG is too high to be of practical use.

If we assume that  $G$  is a spanning tree structure, we can show that the average computation time for BP inference is



$O(m^2 k)$ . Let us assume that any two nodes have an equal chance of being linked by an arc in a spanning tree. As the tree contains  $m^2$  -



**Figure 2:** Constructing a multi layer hidden network. (a) Two subbands, with the coarser subband on top. (b) Procedure creates two wavelet patches, each of which is associated with a subgraph. Subband coefficients are assigned to nodes, as specified in Fig. 1(c). (c) Nodes in the two-layer network are linked by intra-scale (solid) arcs and inter-scale (dashed) arcs. (d) To derive the prior probability, the hidden coefficients are organized into groups a, b, c, and d. In group a, the coefficients are  $\{Pai, ai | i = 1, 2, 3, 4\}$  and they are highlighted in the local blocks of the subbands in (a). Coefficients in the other groups can be derived in a similar manner.

involves  $k^{j-1}$  operations, where  $k$  is the label of a wavelet coefficient. The total cost of the operation for all nodes is

$$\sum_{j=1}^{m \times m-1} k^{j-1} = \frac{k^{m \times m-1} - 1}{k - 1} \quad (1)$$

edges, each node has 1 in-degree at most; and, in total, all the nodes have  $m^2 - 1$  out-degrees. Because all the nodes are treated equally, the average out-degree of each node in a spanning tree is  $\leq 1$ . BP inference in a spanning tree is comprised of two phases: forward message passing and backward message passing, which use the in degree arcs and out-degree arcs respectively to pass Messages to a node. Because the average in-degree and out-degree arc of a node is at most 1, the average number of messaging passing operations on a node is  $O(k)$ . There are  $m^2$  nodes, so the average number of message passing operations is  $O(m^2 k)$  for BP inference in spanning trees. The analysis shows that the cost of BP inference on a spanning tree can be significantly

less than that in a general DAG. However, the power of a spanning tree to express the structure of a wavelet patch is limited. For a subgraph of  $m \times m$  nodes, there are  $m^2 \times m^2$  pairwise relations, but a spanning tree only uses  $m^2 - 1$  of them. Thus, a spanning tree cannot fully capture a complicated structure that contains more than  $m^2 - 1$  highly correlated pairwise relations and a smooth region in which all the nodes are highly correlated to one another.

The joint probability of a spanning tree  $G = (V, E)$  can be formulated by the dependency structure in  $G$  as follows.

Let  $f(v_i | u_i)$  be the probability function associated with arc  $u_i \rightarrow v_i$ , where  $u_i, v_i \in V$ , and let  $u$  be the root of the tree with probability  $f(u)$ . Then, the joint probability of  $G$  is

$$f(G) = f(v_i | u_i) f(u). \quad (2)$$

$$u_i \rightarrow v_i \in E$$

Note that the intra-scale clustering property of wavelet coefficients indicates that the neighboring coefficients in a subband are positively correlated. Thus, the pairwise joint probability can be modeled as a measurement of  $|u_i - v_i|$ , resembling the marginal statistics of the gradient of neighboring nodes with values  $u_i$  and  $v_i$ . We can stack the subgraphs of two adjacent scales to form a two-layer graph. Let  $G_c = (V_c, E_c)$  be the graph corresponding to a subband at the coarser scale of  $G$ . Based on the inter-scale persistence property, inter-scale arcs can be constructed between  $V_c$  and  $V$ , denoted as  $A$ , by combining the parent node and child node (nodes at the same location) in  $G_c$  and  $G$  respectively. Let  $p(v_k)$  in  $V_c$  be the parent node of  $v_k$  in  $V$ ; and let  $u_c$  and  $u$  be the root of the tree in  $G_c$  and  $G$  respectively. Because of the node dependency in the graphs, the joint probability of the resulting directed graph can be formulated as follows:

$$f(G_c, G) = f_c(u_c) \prod_{v_k \rightarrow v_l \in E} f_c(u_j | u_i) \times f(u | p(u)) f(v_l | v_k, p(v_l)) \quad (3)$$

where  $p(u) \rightarrow u \in A$ , and  $p(v_l) \rightarrow v_l \in A$ . Equation (3) gives the joint probability of all the nodes in the graph shown in Fig.2(c). The prior probability of the hidden coefficients in the graph is derived as follows. We take one coefficient from each node to form a data group; for example, Fig. 2 (d) has four data groups.

The joint probability of each data group can be written in the form of Equation (3) by replacing the node variables with the hidden coefficients. Finally, it is assumed that the joint probability of all the data groups is the product of that of each data group.

After constructing the hidden network, we create layer of observation nodes for each noisy wavelet subband. Next, we assign the noisy coefficients to the observation nodes in the same way as the ideal subband coefficients are assigned to the hidden nodes, as shown in Fig.3(a) and (b).

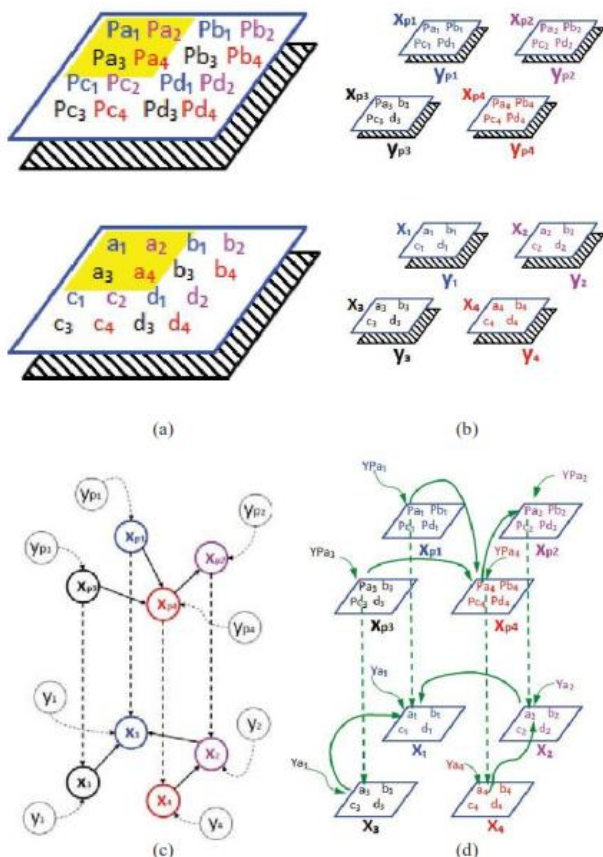
Then, arcs are created to link the observation nodes to the corresponding hidden nodes, as shown in Fig.3(c). BP

inference can now be applied to derive the MAP solution for the denoising problem. Recall that each node in a wavelet patch has  $(m^N)^2$  realizations. If we take one wavelet coefficient from each node and use the BP algorithm to estimate the solution, the denoising process makes  $(mN)^2$  BP inferences.

Note that, as shown in Fig. 3(d), all the wavelet coefficients located in the same rectangular blocks in various subbands are estimated simultaneously by one BP inference.

We calculate the computational cost of image denoising under the proposed model as follows. Because each node has one in-degree and one out-degree inter-scale edge at most, the number of operations that BP inference performs for the multi-layer graph is  $J$  times the number performed for one wavelet patch.

For a spanning tree, each inference involves an average of  $O(m^2 k)$  operations and there are  $(m^N)^2$  BP; thus, the proposed algorithm performs an average of  $O(N^2 k)$  operations for one wavelet patch. As a result, the average BP inference cost of the  $J$  layers network is  $O(JN^2 k)$ .



**Figure 3:** (a) Top layer and the bottom (dashed) layer of each group are composed of hidden subband coefficients and coefficients observed at the corresponding subband, respectively. (b) Coefficients in observed (dashed) layers are organized in the same way as the coefficients in the hidden layers. (c) Observation nodes  $\{y_i, y_j\}$  are created and linked to hidden nodes  $\{x_i, x_j\}$ . (d) For denoising purposes, data in the network is organized into four groups and BP inference is applied to each group. For example,  $i\{ypai, pai\}$ ,

$yai, ai(i = 1, 2, 3, 4)$  is a group. The other three data groups can be derived in a similar manner.

We have explained the basic approach used to construct a spanning tree from a wavelet patch and stack the spanning trees to form a multi-layer network. In addition, we have shown that if the subgraph structure is a spanning tree, the exact MAP solution can be derived efficiently by BP inference. Before describing the construction in detail, we remark that the wavelet coefficients at the same location and scale, but in different orientations, are correlated. However, they are not necessarily positively correlated, so their potential functions are not suitable for modeling as a function of their absolute difference. Therefore, we construct a wavelet Bayesian network for the wavelet subbands in each orientation.

### 3. Constructing Wavelet Bayesian Networks

A Bayesian network, denoted as  $B = (V, E, P)$ , comprises a set of random variables and their conditional dependencies represented by a directed acyclic graph in which the nodes represent the elements in  $V$ . Each edge element in  $E$  takes the form of a directed arc  $x \rightarrow y$ , where  $x$  and  $y \in V$ . The likelihood  $p(y|x) \in P$  of an edge  $x \rightarrow y \in E$  is the conditional probability of observing  $y$  given that  $x$  exists.

We call the Bayesian networks constructed in wavelet domains wavelet Bayesian networks (WBNs). Our primary objective is to construct a WBN from the undecimated discrete wavelet transform (DWT) of a single image. Initially, wavelet decomposition of an image  $F$  yields three images of wavelet coefficients with horizontal, vertical, and diagonal orientations respectively, and one approximate image of  $F$ . Then, at the next scale, the approximated image is further decomposed to obtain three images of the wavelet coefficients and one coarser approximate image of  $F$ .

Let  $W^h_j F(u, v)$ ,  $W^v_j F(u, v)$ , and  $W^d_j F(u, v)$  denote, respectively, the horizontal, vertical, and diagonal images of the wavelet coefficients at scale  $2^j$ ; and let  $A_j$  represent the approximated image at the same scale. If the undecimated DWT is decomposed  $J$  times, we will have wavelet coefficients  $W^h_j F$ ,  $W^v_j F(u, v)$ , and  $W^d_j F$  with  $j = 1, \dots, J$  and the coarsest approximate image  $A_J F$ .

To construct a WBN, we first group subbands with the same orientation together to obtain a horizontal-group, a vertical-group, and a diagonal-group of wavelet coefficients. Then, we construct a Bayesian network  $B$  for each group. Let  $B^h = (V^h, E^h, P^h)$ ,  $B^v = (V^v, E^v, P^v)$ , and  $B^d = (V^d, E^d, P^d)$  denote the Bayesian networks constructed from the horizontal-group, vertical-group, and diagonal-group of wavelet coefficients respectively.

The WBN  $B = (V, E, P)$  is derived from  $B^h, B^v$  and  $B^d$  by,

$$V = V^h \cup V^v \cup V^d \quad (4)$$

$$E = E^h \cup E^v \cup E^d \quad (5)$$

$$P = P^h \cup P^v \cup P^d \quad (6)$$

Next, we explain how to construct the Bayesian network  $B^u(V^u, E^u, P^u)$  that corresponds to the  $u$ -orientation, where  $u \in \{h, v, d\}$ .



### A. Vertex Set $V^u$

Let the size of the input image  $F$  be  $N \times N$ . If  $J$  wavelet decompositions are applied to  $F$ , there will be  $J$  subbands of size  $N \times N$  in each orientation. Let a wavelet patch (matrix of random variable) be of size  $m \times m$ . For each subband, a graph of  $m^2$  variable nodes are formed. We then associate each random variable in the wavelet patch to a variable node in the graph. Without loss of generality, we assume that  $m$  divides  $N$ . Each subband can then be partitioned into  $(m^N)^2$  blocks, each of size  $m \times m$ . Then,  $(m^N)^2$  wavelet coefficients sampled from the subband are assigned to each variable node.

Let  $x^u(i, k)$ , with  $j = 1, \dots, J$  and  $i, k = 0, \dots, m-1$ , denote the  $(i, k)$  variable node in the  $j$ -th subband. In our  $j$  construction, the  $(m^N)^2$  wavelet coefficients assigned to node  $x^u(i, k)$  are sampled from  $W^u_j F(i+mp, k+mq)$  with  $p, q = 0, \dots, m^N-1$ . If we represent each node as a vertex in the Bayesian network  $B^u$  will be  $V^u = \{x^u_j(i, k) | i, k = 0, \dots, m-1; j = 1, \dots, J\}$  (7)

$$V^u = \{x^u_j(i, k) | i, k = 0, \dots, m-1; j = 1, \dots, J\} \quad (7)$$

and the  $(m^N)^2$  wavelet coefficients can be regarded as sampled from some unknown distribution of a random variable. Figs. 1(a), (b), and (c) show the procedure used to construct the vertex set for a subband of  $4 \times 4$  coefficients.

### B. Edge Set $E^u$

The arcs (directed edges) in  $B^u$  can be divided into two disjoint sets,  $E_o^u$  and  $E_i^u$ , where  $E_o^u$  comprises the (interscale) edges incident to vertices at different scales, and  $E$  comprises the (intra-scale) edges incident to vertices at the same scale. The persistence property of the wavelet transforms indicates that large/small values of wavelet coefficients tend to occur at the same spatial locations in subbands at adjacent scales. The property can be used to construct arcs in  $E_o^u$  by linking a vertex at the coarser scale  $j+1$  to the vertex of the same index at the finer scale  $j$ ; that is,

$$E_o^u = \{x^{u,j+1}(i, k) \rightarrow x^u_j(i, k) | i, k = 0, \dots, m-1\}$$

$$\text{and } j = 1, \dots, J-1. \quad (8)$$

The edges in  $E_i^u$  represent the connections between vertices at the same scale and orientation. Constructing the edges corresponds to deriving the Bayesian network on the nodes  $x^u_j(i, k)$  that best represent the joint probability of the nodes at the same scale  $j$  and orientation  $u$ . However, as discussed in Section II, BP inference is computationally intractable if the Bayesian network is a general graph. Thus, we limit the solution space to spanning trees so that we can derive an efficient solution by using the *maximal weighted spanning tree* (MWST) algorithm [33]–[36]. A maximal weighted spanning tree is a spanning tree whose weight is greater than or equal to the total weight of every other spanning tree. The optimum weighted spanning tree can be derived by minimizing the relative entropy (Kullback-Leibler distance)  $D(p||q)$  between the probability functions  $p$  and  $q$ .

In the following, we show how the spanning tree that minimizes  $D(p||q)$  is equivalent to the tree that maximizes the weighted summation of conditional mutual information on all the edges of the tree. Let  $\mathbf{x}$  be the vector of variables  $x_1, \dots, x_n$ ; let  $p(i)$  and  $b(i)$  denote the indices of the parent nodes and the sibling nodes of  $x_i$  respectively; and let  $q$  be the induced probability of the spanning tree. Then, we have

$$q(\mathbf{x}) = \prod_{j=1}^n p(x_j | x_{p(j)}, x_{b(j)}). \quad (9)$$

Note that nodes at the coarsest scale do not have parent nodes. To find the optimal spanning tree, we minimize the relative entropy between  $q(\mathbf{x})$  and the joint probability  $p(\mathbf{x})$  as follows:

$$\begin{aligned} D(p||q) &= \sum_{\mathbf{x}} p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})} \\ &= \sum_{\mathbf{x}} p(\mathbf{x}) \log p(\mathbf{x}) - \sum_{\mathbf{x}} p(\mathbf{x}) \log q(\mathbf{x}). \end{aligned} \quad (10)$$

Since minimizing  $D(p||q)$  is equivalent to maximizing  $\sum_{\mathbf{x}} p(\mathbf{x}) \log q(\mathbf{x})$ , we can derive the following:

$$\begin{aligned} \sum_{\mathbf{x}} p(\mathbf{x}) \log q(\mathbf{x}) &= \sum_{\mathbf{x}} p(\mathbf{x}) \log \prod_j p(x_j | x_{p(j)}, x_{b(j)}) \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)})}{p(x_{p(j)})} \\ &\quad + \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)})}{p(x_{p(j)})} \\ &\quad + \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}, x_{b(j)})} - \log \frac{p(x_j, x_{p(j)})}{p(x_{p(j)})}. \end{aligned} \quad (11)$$

The second term in the last line of the above equation can be re-written as follows:

$$\begin{aligned} \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}, x_{b(j)})} &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)}) \cdot p(x_{p(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \\ &= \sum_{j, b(j) \neq \emptyset} p(\mathbf{x}) \log \frac{p(x_j, x_{p(j)}, x_{b(j)})}{p(x_{p(j)}) \cdot p(x_j, x_{b(j)})} \end{aligned} \quad (12)$$

**Algorithm 1** Kruskal's Algorithm

- 1) Calculate the  $m^2 \times m^2$  weights of the arcs between any two nodes. Note that the weight of arc  $x \rightarrow y$  may be different from that of arc  $y \rightarrow x$ .
- 2) Sort the weights of the  $m^2 \times m^2$  arcs and compile a list in non increasing order so that the weight  $w_i$  is not less than the weight  $w$  with  $i < j$ .
- 3) Assume that the initial spanning tree is empty and add the arc of weight  $w_1$  to the empty tree.
- 4) Do until the tree is a spanning tree:
  - a) Let  $w_i$  be the next weight on the list.
  - b) Check if adding the arc of weight  $w_i$  to the tree creates a cycle.
- 5) End

and the sample probabilities are calculated as follows:

$$\begin{aligned} f(x_i = u) &= b^f(x_i = u, x_j = v) \\ f(x_j = v) &= a^f(x_i = a, x_j = v) \end{aligned} \quad (15)$$

It is well-known that the joint frequencies  $f(x_i = u, x_j = v)$ ,  $f(x_i = u)$  and  $f(x_j = v)$  are maximum-likelihood estimators for the probabilities  $p(x_i = u, x_j = v)$ ,  $p(x_i = u)$  and  $p(x_j = v)$  respectively. Therefore, the sample mutual information can be computed as

$$I(x_i, x_j) = \sum_{u,v} f(x_i = u, x_j = v) \log \frac{f(x_i = u, x_j = v)}{f(x_i = u) f(x_j = v)} \quad (16)$$

Then, we use  $I(x_i, x_j)$  instead of  $I(x_i, x_j | j = v)$  when calculating the weights in the above algorithm. For each scale  $j \in \{1, \dots, J\}$ , we execute Kruskal's algorithm to obtain a contains all the edges in all spanning tree. The edge set  $E_i^u$  the spanning trees.

Figure(d) shows an example of a network with intra-scale edges derived by the MWST algorithm from the four nodes in Fig. 1(c); and Fig. 2(c) shows an example of a multilayer network where the inter-scale and intra-scale edges are constructed from the nodes in Fig. 2(b). The WBN  $B$  in Fig. 4 is comprised of three oriented Bayesian networks,  $B^h, B^v$  and  $B^d$ .

**C. Probability Model  $P^u$**

There are two types of arcs in a Bayesian network  $B^u$ : 1) the inter-scale parent-child arc, which connects a node with its coarser-scale parent; and 2) the intra-scale sibling arc, which connects two nodes of the same scale. To obtain the probability inference, we need to model the probability function on each arc.

Simoncelli [22] exploited the persistence property of wavelet transforms and proposed a joint statistical model of a "child" coefficient conditioned on the coarse-scale "parent" coefficients at the same spatial locations in all orientations.

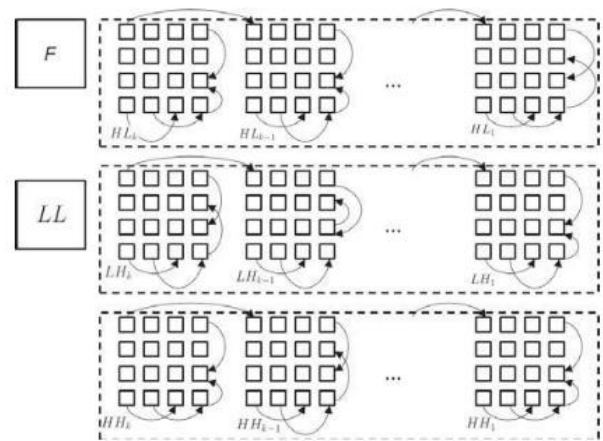


Fig. 4. WBN  $B$  has three components,  $B^h, B^v$ , and  $B^d$ , derived from the horizontal, vertical, and diagonal orientation wavelet coefficients, respectively, as shown from top to bottom.  $F$  and  $LL$  are, respectively, the input and the coarsest approximation of  $F$ .  $k$  is the scale index.

Let  $\{x_{pk}\}$  comprise the parent coefficients of the child coefficient  $x$ . Then, the probability function of  $x$  conditioned on  $\{x_{pk}\}$  is defined as the following Markov model:

$$p(x | \{x_{pk}\}) = N(0; 2w_k x_{pk}^2 + \alpha) \quad (17)$$

which is a conditional Gaussian density function whose variance is a linear function of the squared coefficients in  $\{x_{pk}\}$  with weights  $\{w_k\}$ , and  $\alpha$  is the bias of the variance. We simplify Simoncelli's joint probability model to represent the  $k$ pdf of a parent-child arc in  $E_o^u$ . Specifically, we remove the dependency of the coarse-scale parents' coefficients in all orientations except that of the child coefficient and set the bias of variance  $\alpha$  to zero. Thus, there is only one parent coefficient  $x_p$  for  $x$ . Then, the joint probability of the parent-child arc in  $E_o^u$  is modeled as

$$f_o^u(x | x_p) = N(0; 2wx_p^2) \quad (18)$$

where  $w$  is a chosen parameter.

Other researchers have observed that large/small wavelet coefficients of the same sign tend to occur near each other in a subband. The observations are summarized as the clustering property of the wavelet transform. The property means that, in a subband, there is a strong positive correlation between wavelet coefficients that are in close proximity and dependent on each other. The ROF approach [28] exploits the clustering property between neighboring pixels by using TV norms, which are essentially  $L_1$  norms of the derivative of the values of two neighboring coefficients, to estimate the discontinuities in an image. Equivalently, the joint probability function between two coefficients in close proximity is modeled as the Laplacian distribution of the absolute of the difference of two wavelet coefficients [42]. In the construction of intra-scale edges, Kruskal's algorithm selects the arc  $x \rightarrow z \in E_i^u$ , where the mutual information between  $x$  and  $z$  is high; that is,  $x$  and  $z$  are highly correlated. Thus, we utilize a similar concept to model the probability of  $z$  conditioned on  $x$  as the following Laplacian distribution function:

$$f_i^u(z | x) \propto \lambda \exp(-\lambda |x - z|) \quad (19)$$

where  $\lambda$  is the scale parameter of the Laplacian distribution, and  $\lambda |x - z|$  is the potential function of the local  $1 \times 1$  clique.

#### 4. Wavelet Bayesian Networks For Denoising

In this section, we consider using the wavelet Bayesian network to model the prior probability of the original image for the image denoising problem, which involves removing white and homogenous Gaussian noise with zero mean and known variance from an image. To infer the probability for denoising, we associate each variable node  $x$  in Bayesian network  $B$  with an observation node  $y$  and create the arc  $y \rightarrow x$ . The probability function of  $x$  conditioned on the observed value of  $y$  is modeled as

$$f(x|y) = \frac{1}{\sigma_n^2} \exp\left(-\frac{(x-y)^2}{2\sigma_n^2}\right) \quad (20)$$

where  $\sigma_n^2$  is the variance of the zero mean Gaussian white noise and  $\rho$  depends on the scale and wavelets.

Recall that the variable nodes in the Bayesian network  $B$  are represented by  $\{x^u_j(i, k) | u \in \{h, v, d\}; j = 1, \dots, J; i, k = 0, \dots, m-1\}$  (Equation (7)). Let  $y^u_j(i, k)$  denote the observation node corresponding to  $x^u_j(i, k)$ , and  $P_n$  denote the collections of  $\{y(i, k)\}$ ; and let  $Y = \{y^u_j(i, k)\}$ , the arcs  $\{y^u_j(i, k) \rightarrow x(i, k)\}$ , and the probability functions  $\{f(x^u_j(i, k))\}$  respectively. The WBN  $B_n(i, k) | y^u_j$  for image denoising is constructed and represented as  $B_n = (V \cup Y, E \cup E_n, P \cup P_n)$ .

Let a noisy image  $Z = F + N$ , where  $F$  is the original image and  $N$  is zero-mean white Gaussian noise. As shown by the simple example in Fig. 3(d), each wavelet coefficient of  $Z$  is assigned to one observation node in  $B_n$ . That is, the coefficient  $W^u_j Z(i+mp, k+mq)$  is assigned to observation node  $y^u_j(i+k)$ , where  $p, q = 0, \dots, m^N - 1$ ; thus, each observation node has  $(m^N)^2$  observation values and realizations. If we take one wavelet coefficient from each observation node, we can obtain (realizations of  $B_n$ , denoted as  $(p, q)$  with  $p, q = 0, \dots, m^N - 1$ ). Note that, similar to the example shown in Fig. 3(d), the wavelet coefficients assigned to the  $(p, q)$ -th realization are taken locally from the  $(i, k)$ -blocks of all subbands at scale  $2^j$  and orientation  $u$ .

We use the message passing algorithm to obtain the estimated wavelet coefficients of each realization. First, we convert WBN  $B_n$  to a factor graph  $F_n$ , and then use the max product algorithm to derive the estimated wavelet coefficients. The conversion of  $B_n$  to  $F_n$  and the max-product message passing schemes are standard techniques. For completeness, we provide them in Appendix A. The last step of the maxproduct algorithm calculates the marginal probability of each  $v$  in  $F_n$ . Let  $N(x)$  represent the neighboring factor nodes of variable node  $x$  in  $F_n$  variable node  $V$  in  $F_n$ . In addition, let  $x_p$  and  $x_c$  denote, respectively, the parent variable node and child variable node of  $x$  in  $B_n$ ; and let  $\{x_s\}$  denote the sibling variable nodes of  $x$  in  $B_n$ . The value of  $x^*$  can be estimated based on whether  $x$  has a child node.

Case 1:  $x$  has a child node  $x_c$ .

$$x^* = \arg \max_{x \in N(x)} \frac{1}{x} \exp(-J_c(x)) \quad (21)$$

where

$$J(x) = \frac{(x-y)^2}{2\sigma_y^2} + \frac{x^2}{2\sigma_x^2} + \frac{x_c^2}{2\sigma_c^2} + \lambda \sum_j |x - x_j| \quad (22)$$

In Equation (22),  $\rho = \rho(y, x_p, x_c, \{x_j\})$  is independent of  $x$ , is the variance of the wavelet coefficients associated with observation node  $y$ . The variance  $\sigma^2$  can be written as  $\sigma^2 \rho$ , where  $\rho$  depends on the scale and the wavelets. In Appendix B, we show that  $\rho = 1$  if the wavelets are orthogonal. Let

$$K(x) = \frac{1}{x} \exp(-J_c(x)); \quad (23)$$

then

$$\frac{d}{dx} \log K(x) = \frac{d}{dx} \log \frac{1}{x} - \frac{d}{dx} J_c(x) = -\frac{1}{x} - \frac{d}{dx} J_c(x) \quad (24)$$

According to Equation (22)

$$J(x) = \frac{(x-y)^2}{2\sigma_y^2} + \frac{x^2}{2\sigma_x^2} + \frac{x_c^2}{2\sigma_c^2} + \lambda \sum_j \text{sign}(x - x_j) \quad (25)$$

Where  $\text{sign}(x)$  returns the sign of  $x$  if  $x \neq 0$ ; otherwise, it returns 0. Hence

$$\frac{1}{x} - \frac{d}{dx} J(x) = \frac{x-y}{\sigma_y^2} + \frac{1}{x} - \frac{x}{\sigma_x^2} - \frac{x_c}{\sigma_c^2} + \lambda \sum_j \text{sign}(x - x_j) \quad (26)$$

$$y - \lambda \sigma_y^2 \sum_j \text{sign}(x_{t-1} - x_j) \pm \sqrt{\frac{y - \lambda \sigma_y^2 \sum_j \text{sign}(x_{t-1} - x_j)^2}{4\sigma_y^2 \frac{x_{t-1}^2}{x_p^2} - \frac{x_c^2}{x^2} + \omega}} \quad (27)$$

Case 2:  $x$  does not have a child node ( $x$  is a node at the finest wavelet scale). We can set  $x_c = 0$  in Equation (22) and obtain

$$J(x) = \frac{(x-y)^2}{2\sigma_y^2} + \frac{x^2}{2\sigma_x^2} + \lambda \sum_j |x - x_j| \quad (28)$$

and then

$$J(x) = \frac{(x-y)}{\sigma_y^2} + \frac{x}{\sigma_x^2} + \lambda \sum_j \text{sign}(x - x_j) \quad (29)$$

#### Algorithm 2 WBN Denoising Algorithm

- 1) Wavelet representation: calculate the undecimated DWT of an  $N \times N$  noisy image  $Z$  to obtain the horizontal, vertical, and diagonal subbands  $\{W^u_j Z | j=1, \dots, J\}$ , vertical, and diagonal subbands  $\{W^u_j Z | j=1, \dots, J\}$  and  $\{W^u_j Z | j=1, \dots, J\}$ , respectively, as well as the coarsest approximate image  $A_J Z$ .
- 2) Create a vertex set  $V^u$ : let the parameter  $m$  divide  $\text{Foreach subband } W^u_j Z$ , create  $m^2$  variable nodes  $\{x^u_j(i, k) | i, k=0, \dots, m-1\}$ ; then assign  $(m^N)^2$  wavelet coefficient.
- 3) Create sibling edges ( $E^u_j$ ) for nodes at the same scale and orientation: derive the empirical probability of each variable from the frequency counts of  $(mN/2)$  wavelet coefficients assigned to the node; then, use Kruskal's algorithm to derive the maximal weighted spanning tree from the  $m^2$  nodes in each subband.
- 4) The root of  $J(x) = 0$  can be derived by the following fixed point algorithm. Let  $x$  be the estimated value after the  $t$ -th iteration. Then,  $x_{t+1}$  can be derived from  $x_t$  as follows:



$$x_t = 1 + \frac{\sigma_y^2 - 1}{\sigma_p^2} y - \sigma_y \text{sign}(x_t - 1 - x_j). \quad (30)$$

The iterative estimation stops when the difference between the values  $x_{t-1}$  and  $x_t$  is smaller than some given threshold.

## 5. Denoising Algorithm and Experimental Results

In this section, we present our denoising algorithm, explain its implementation, and compare its performance with that of other methods. The proposed algorithm is summarized in Algorithm 1. The steps are as follows: Step (1) calculates the undecimated DWT of the input image; Steps (2) to (5) construct the WBN  $B$ ; and Steps (6) to (8) create the WBN  $B_n$  used for denoising purposes; In Step (9), the wavelet coefficients are estimated from  $B_n$  by applying the max-product algorithm to the factor graph  $F$  for each realization of  $B_n$ . We use CDF 9/7 filters to process the undecimated DWT. Because CDF 9/7 filters are close to orthogonal wavelet filters, the noise variance of subbands at all scales can be set at  $\sigma_n^2$ , which is the image noise variance (see Appendix B). The variance  $\sigma_n^2$  is used in the Wiener filtering in Step (2) as well as in deriving the MAP estimation of the wavelet coefficients in Step (9). The frequency count in Step (4) indicates the number of wavelet coefficients in a quantization bin. The size of a subband's quantization bin is set at 14 of the standard deviation, measured from the wavelet coefficients in the subband.

We conduct experiments on two sets of images. Set I contains nine gray scale images (size  $512 \times 512$  or  $256 \times 256$ ) downloaded from the USC-SIPI image database [43]; and Set II (shown in Fig. 6) contains nine gray scale textures, some of which are from the Brodatz texture set. The parameter settings of the WBN denoising algorithm evaluated in the experiments are:  $J = 4$  (the number of wavelet decompositions),  $\omega = 0.64$  (Equation (18)),  $\lambda = 0.45$  (Equation (19)), and the parameter  $m = 4$  in Step (3). Each subband represents a  $512 \times 512$  image and contains  $4 \times 4$  nodes. The WBN  $B$  has  $16 \times 4 \times 3$  variable nodes because there are four subbands in each of the three scales, half of which are variable nodes and the rest are observation orientations. Hence, WBN  $B_n$ , we assign  $128 \times 128$  wavelet coefficients. Note that this number is large enough to nodes. For each observation node in  $B_n$ , derive the empirical probability in Step (4) of the WBN algorithm. There are also  $128 \times 128$  realizations of  $B_n$ .

Next, we present the experimental results derived by our algorithm and compare its performance with that of two state-of-the-art algorithms: the BM3D algorithm [13] and the BLS-GSM algorithm [23]. The source codes for the BLS-GSM and BM3D methods are available from the websites of the respective authors. Our algorithm is executed on an Intel Core2Quad Q9300 CPU, with Windows XP and Matlab 2007a. Tables I and II list the averages of five denoised PSNR results of the three compared methods for the images in Sets I and II respectively; white noise was added with  $\sigma_n^2 = 10, 15, \dots, 35$ . In the tables, we divide the images into two groups and compare the average PSNR gain of our method over the other methods on the images in each group.

Table III summarizes the results in Tables I and II. Note that the Figure print and Baboon images are regarded as texture images because they are dominated by texture information. Interestingly, the PSNR gain of our method over BM3D is 0.1 dB for texture images (Group 2), but only 0.02 dB for the other images (Group 1). Moreover, from Table III, we observe that the gain of BM3D

over BLS-GSM is smaller on texture images (0.06 dB) than on the other images (1.0 dB). Texture images contain rich across-scales information. Since both BLS-GSM and our method exploit inter-scale coefficients for data estimation, they are more adaptive to texture images than non-texture images.

TABLE I AVERAGE PSNR RESULTS FOR IMAGE SET I

Image	Noise Level	PSNR		
		BLS-GSM	BM3D	Ours
<i>Einstein</i> 512 × 512	$\sigma_n = 10$	34.1406	34.4392	34.4837
	$\sigma_n = 15$	32.6818	33.0331	33.0874
	$\sigma_n = 20$	31.7412	32.1694	32.2287
	$\sigma_n = 25$	31.0201	31.4186	31.4682
	$\sigma_n = 30$	30.4306	30.8709	30.9175
	$\sigma_n = 35$	29.9372	30.3777	30.4116
<i>Boat</i> 512 × 512	$\sigma_n = 10$	33.5080	33.8883	33.9127
	$\sigma_n = 15$	31.6452	32.1067	32.1375
	$\sigma_n = 20$	30.3194	30.8554	30.8872
	$\sigma_n = 25$	29.2846	29.8356	29.8612
	$\sigma_n = 30$	28.4395	29.0954	29.1247
	$\sigma_n = 35$	27.7301	28.2992	28.3286
<i>Barbara</i> 512 × 512	$\sigma_n = 10$	33.1518	34.9567	34.9654
	$\sigma_n = 15$	30.7724	33.0666	33.0735
	$\sigma_n = 20$	29.0984	31.7376	31.7546
	$\sigma_n = 25$	27.8214	30.7176	30.7354
	$\sigma_n = 30$	26.7998	29.7049	29.7116
	$\sigma_n = 35$	25.9775	28.8879	28.6753
<i>Lena</i> 512 × 512	$\sigma_n = 10$	35.9226	36.6367	36.6375
	$\sigma_n = 15$	34.1105	34.8782	34.8886
	$\sigma_n = 20$	32.8047	33.0567	33.1246
	$\sigma_n = 25$	31.7891	32.5501	32.4874
	$\sigma_n = 30$	30.9633	31.6531	31.5637
	$\sigma_n = 35$	30.2780	31.0301	30.9353
<i>Camerman</i> 256 × 256	$\sigma_n = 10$	33.3739	34.1355	34.1624
	$\sigma_n = 15$	31.0334	31.8449	31.8702
	$\sigma_n = 20$	29.5762	30.3797	30.4054
	$\sigma_n = 25$	28.5242	29.4118	29.4837
	$\sigma_n = 30$	27.6872	28.5516	28.6476
	$\sigma_n = 35$	26.9689	27.8758	27.9645
<i>House</i> 256 × 256	$\sigma_n = 10$	35.3223	36.6638	36.7135
	$\sigma_n = 15$	33.7257	34.9028	34.9476
	$\sigma_n = 20$	32.5237	33.7349	33.7782
	$\sigma_n = 25$	31.5753	32.9084	32.9406
	$\sigma_n = 30$	30.7650	32.1240	32.1574
	$\sigma_n = 35$	30.0579	31.5103	31.5385
<i>Pepper</i> 256 × 256	$\sigma_n = 10$	34.0542	34.6889	34.7184
	$\sigma_n = 15$	32.0136	32.7290	32.7602
	$\sigma_n = 20$	30.5676	31.2671	31.2957
	$\sigma_n = 25$	29.4553	30.2223	30.2607
	$\sigma_n = 30$	28.5640	29.3090	29.3276
	$\sigma_n = 35$	27.8161	28.5795	28.6014
<i>Fingerprint</i> 512 × 512	$\sigma_n = 10$	32.1994	32.4628	32.7284
	$\sigma_n = 15$	29.9283	30.2822	30.5975
	$\sigma_n = 20$	28.3391	28.8085	29.0732
	$\sigma_n = 25$	27.0946	27.7112	28.0364
	$\sigma_n = 30$	26.0682	26.8245	27.1262
	$\sigma_n = 35$	25.1981	26.0995	26.4075
<i>Baboon</i> 512 × 512	$\sigma_n = 10$	30.3735	30.5798	30.6623
	$\sigma_n = 15$	27.8549	28.1390	28.2505
	$\sigma_n = 20$	26.1750	26.5307	26.6275
	$\sigma_n = 25$	24.9483	25.3495	25.4374
	$\sigma_n = 30$	24.0063	24.4407	24.5238
	$\sigma_n = 35$	23.2643	23.6554	23.7522

TABLE II  
AVERAGE PSNR RESULTS FOR IMAGE SET II

Image	Noise Level	PSNR		
		BLS-GSM	BM3D	Our
<i>F0</i> 512 × 512	$\sigma_n = 10$	31.6031	32.52	32.6212
	$\sigma_n = 15$	29.6083	30.3213	30.4143
	$\sigma_n = 20$	28.2910	28.8595	28.9314
	$\sigma_n = 25$	27.2514	27.7279	27.8087
	$\sigma_n = 30$	26.3668	26.7751	26.8644
	$\sigma_n = 35$	25.6029	25.9243	26.0136
<i>F7</i> 512 × 512	$\sigma_n = 10$	29.9444	29.8987	30.1253
	$\sigma_n = 15$	27.2804	27.3043	27.3765
	$\sigma_n = 20$	25.5299	25.6216	25.6782
	$\sigma_n = 25$	24.2851	24.3522	24.4186
	$\sigma_n = 30$	23.3451	23.401	23.4862
	$\sigma_n = 35$	22.602	22.5878	22.6574
<i>p3</i> 512 × 512	$\sigma_n = 10$	31.1789	31.3155	31.3385
	$\sigma_n = 15$	28.8139	29.0087	29.0274
	$\sigma_n = 20$	27.2571	27.4678	27.4936
	$\sigma_n = 25$	26.1114	26.3278	26.3414
	$\sigma_n = 30$	25.2116	25.4362	25.4535
	$\sigma_n = 35$	24.4756	24.7034	24.7153
<i>F3</i> 512 × 512	$\sigma_n = 10$	29.2042	29.3575	29.5361
	$\sigma_n = 15$	26.3055	26.3832	26.5738
	$\sigma_n = 20$	24.3449	24.354	24.5363
	$\sigma_n = 25$	22.8854	22.8154	23.0236
	$\sigma_n = 30$	21.7442	21.6195	21.8514
	$\sigma_n = 35$	20.8213	20.6102	20.9144
<i>w2</i> 512 × 512	$\sigma_n = 10$	29.0994	29.4042	29.496
	$\sigma_n = 15$	26.499	26.798	26.8816
	$\sigma_n = 20$	24.8784	25.1843	25.1547
	$\sigma_n = 25$	23.7188	24.0512	24.1376
	$\sigma_n = 30$	22.8012	23.1637	23.2473
	$\sigma_n = 35$	22.0279	22.3965	22.5062
<i>w6</i> 512 × 512	$\sigma_n = 10$	29.6007	30.0207	30.1433
	$\sigma_n = 15$	27.1382	27.5495	27.6887
	$\sigma_n = 20$	25.5041	25.964	26.1186
	$\sigma_n = 25$	24.2672	24.8184	24.9782
	$\sigma_n = 30$	23.2643	23.9149	24.1045
	$\sigma_n = 35$	22.4208	23.1486	23.3228
<i>F6</i> 512 × 512	$\sigma_n = 10$	28.5821	28.5956	28.7025
	$\sigma_n = 15$	25.4728	25.5213	25.6175
	$\sigma_n = 20$	23.3581	23.4603	23.563
	$\sigma_n = 25$	21.7955	21.9042	22.0154
	$\sigma_n = 30$	20.6055	20.6784	20.8706
	$\sigma_n = 35$	19.6868	19.6259	19.8154
<i>F8</i> 512 × 512	$\sigma_n = 10$	29.0382	28.9781	29.1012
	$\sigma_n = 15$	26.1317	26.1037	26.2387
	$\sigma_n = 20$	24.2006	24.2121	24.3382
	$\sigma_n = 25$	22.7888	22.8208	22.9647
	$\sigma_n = 30$	21.7053	21.734	21.883
	$\sigma_n = 35$	20.8487	20.8181	21.0225
<i>S2</i> 512 × 512	$\sigma_n = 10$	29.3274	29.3933	29.4072
	$\sigma_n = 15$	26.5777	26.5264	26.5942
	$\sigma_n = 20$	24.895	24.6771	24.9125
	$\sigma_n = 25$	23.8043	23.4976	23.8315
	$\sigma_n = 30$	23.059	22.7346	23.0627
	$\sigma_n = 35$	22.523	22.2268	22.5434

In Figures 5, 6, and 7, we perceptually compare some images that were denoised by the three methods. The perceptual quality of the images denoised by BM3D and our method is better than that of the BLS-GSM images. The highlighted regions in the figures compare certain details of the denoised images derived by all three methods with the corresponding

Table III: Summary of Our PSNR Gain

Image Group	Average PSNR Gain	
	versus BLS-GSM	versus BM3D
Group 1 (First seven images in Table I)	1.0239	0.0207
Group 2 (Fingerprint + Baboon + Images in Table II)	0.2922	0.1268
All Images in Tables I and II	0.9403	0.0594



(a)



(b)



(c)



(d)

**Figure 5:** Comparison of the denoised images derived by BLS-GSM, BM3D, and our algorithm. The noise standard deviation is  $\sigma = 25$ . (a) Original *Lena* image. (b) Denoised result of the BLS-GSM algorithm. (c) Denoised result of the BM3D algorithm. (d) Denoised result of our algorithm.

## 6. Conclusion

The Bayesian formula indicates that the denoising problem is essentially a prior probability modeling and estimation task. In this paper, we present constructive data-adaptive procedure that derives a hidden graph structure from the wavelet coefficients. The graph is then used to model the prior probability of the original image for denoising purposes. Moreover, we show that if the network is a spanning tree, the standard BP algorithm can estimate MAP



efficiently. We compare our denoised results with those derived by other approaches, including BM3D, and demonstrate that our method yields a better PSNR and better perceptual quality on the textured areas of an image. Extending our method to content sensitive wavelet patches is an issue that merits future study. We will also investigate ways to speed up our algorithm's execution time.

## 7. Acknowledgment

The authors would like to thank the reviewers for their insightful comments, which have helped to significantly improve the quality of this paper.

## References

- [1] M. J. Jordan, *Learning in Graphical Models*. Cambridge, MA: MIT Press, 1998.
- [2] G. F. Cooper and E. Herskovits, "A Bayesian method for the induction of probabilistic networks from data," *Mach. Learn.*, vol. 9, no. 4, pp. 309–347, Oct. 1992.
- [3] D. Heckerman, D. Geiger, and D. M. Chickering, "Learning Bayesian networks: The combination of knowledge and statistical data," *Mach. Learn.*, vol. 20, no. 3, pp. 197–243, 1995.
- [4] D. Heckerman, "A tutorial on learning Bayesian networks," Microsoft Research, Mountain View, CA, Tech. Rep. MSR-TR-95-06, 1995.
- [5] D. M. Chickering, D. Heckerman, and C. Meek, "A Bayesian approach to learning Bayesian networks with local structure," in *Proc. 13th Conf. Uncertainty Artif. Intell.*, 1997, pp. 80–89.
- [6] T. Hastie, R. Tibshirani, and J. Friedman, *The Elements of Statistical Learning*. New York: Springer-Verlag, 2001.
- [7] M. S. Crouse, R. D. Nowak, and R. G. Baraniuk, "Wavelet -based statistical signal processing using hidden Markov models," *IEEE Trans. Signal Process.*, vol. 46, no. 4, pp. 886–902, Apr. 1998.
- [8] G. Chang, B. Yu, and M. Baraniuk, "Spatially adaptive wavelet thresholding with context modeling for image denoising," *IEEE Trans. Image Process.*, vol. 9, no. 9, pp. 1522–1531, Sep. 2000.
- [9] J. K. Romberg, H. Choi, and R. G. Baraniuk, "Bayesian tree-structured image modeling using wavelet-domain hidden Markov models," *IEEE Trans. Image Process.*, vol. 10, no. 7, pp. 1056–1068, Jul. 2001.
- [10] P. Chatterjee and P. Milanfar, "Is denoising dead?" *IEEE Trans. Image Process.*, vol. 19, no. 4, pp. 895–911, Apr. 2010.
- [11] A. Buades, B. Coll, and J. M. Morel, "A review of image denoising algorithms, with a new one," *Multiscale Model. Simul.*, vol. 4, no. 2, pp. 490–530, 2005.
- [12] C. Kervrann and J. Boulanger, "Optimal spatial adaptation for patch-based image denoising," *IEEE Trans. Image Process.*, vol. 15, no. 10, pp. 2866–2878, Oct. 2006.
- [13] K. Dabov, A. Foi, V. Katkovnik, and K. Egiazarian, "Image denoising by sparse 3-D transform-domain collaborative filtering," *IEEE Trans. Image Process.*, vol. 16, no. 8, pp. 2080–2095, Aug. 2007.
- [14] P. Chatterjee and P. Milanfar, "Patch-based near-optimal image denoising," *IEEE Trans. Image Process.*, vol. 21, no. 4, pp. 1635–1649, Apr. 2011.
- [15] M. Elad and M. Aharon, "Image denoising via sparse and redundant representations over learned dictionaries," *IEEE Trans. Image Process.*, vol. 15, no. 12, pp. 3736–3745, Dec. 2006.
- [16] J. Mairal, F. Bach, J. Ponce, and G. Sapiro, "Non-local sparse models for image restoration," in *Proc. IEEE Int. Conf. Comput. Vis.*, Kyoto, Japan, 2009, pp. 2272–2279.
- [17] P. Milanfar, "A tour of modern image filtering: New Insights and methods, both practical and theoretical," *IEEE Signal Process. Mag.*, vol. 30, no. 1, pp. 106–128, Jan. 2013.
- [18] M. Malfait and D. Roose, "Wavelet-based image denoising using a Markov random field a priori model," *IEEE Trans. Image Process.*, vol. 6, no. 4, pp. 549–565, Apr. 1997.
- [19] S. Mallat and W. L. Hwang, "Singularity detection and processing with wavelets," *IEEE Trans. Inf. Process.*, vol. 38, no. 2, pp. 617–643, Mar. 1992.
- [20] C. L. Tu, W. L. Hwang, and J. Ho, "Analysis of singularities from modulus maxima of complex wavelets," *IEEE Trans. Inf. Process.*, vol. 51, no. 3, pp. 1049–1062, Mar. 2005.